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Cluster randomized clinical trials in orthodontics: design, analysis and reporting issues

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SUMMARY Cluster randomized trials (CRTs) use as the unit of randomization clusters, which are usually defined as a collection of individuals sharing some common characteristics. Common examples of clusters include entire dental practices, hospitals, schools, school classes, villages, and towns. Additionally, several measurements (repeated measurements) taken on the same individual at different time points are also considered to be clusters. In dentistry, CRTs are applicable as patients may be treated as clusters containing several individual teeth. CRTs require certain methodological procedures during sample calculation, randomization, data analysis, and reporting, which are often ignored in dental research publications. In general, due to similarity of the observations within clusters, each individual within a cluster provides less information compared with an individual in a non-clustered trial. Therefore, clustered designs require larger sample sizes compared with non-clustered randomized designs, and special statistical analyses that account for the fact that observations within clusters are correlated. It is the purpose of this article to highlight with relevant examples the important methodological characteristics of cluster randomized designs as they may be applied in orthodontics and to explain the problems that may arise if clustered observations are erroneously treated and analysed as independent (non-clustered).

Introduction

Clinical trials in dentistry aim to prospectively evaluate which treatment modalities are most beneficial to patients. Some common research scenarios are the survival of various types of implants (Tellemann *et al.*, 2011), the effectiveness of different interventions for guided tissue regeneration for periodontal infra-bony defects (Needleman *et al.*, 2006), effectiveness of fluoride mouthrinses (Ganss *et al.*, 2010) and interventions that compare risk of failure for different etching methods used in orthodontic bonding (Pandis *et al.*, 2006). Depending on the unit of randomization, clinical trials can be described as individually or cluster randomized trials (CRTs). Clinical trials where groups of individuals or clusters are randomized to receive the same treatment are known as CRTs. Clusters may be aggregates of individuals such as schools or communities, or even multiple measurements on the same person.

There are a number of reasons why a cluster randomized design would be preferred to an individually randomized design:

1. Trial management: From a design perspective, it may be logistically preferable to randomize to clusters consisting of a few, several or many persons who share some common characteristics (Altman and Bland, 1997; Wang and Bakhai, 2006; Hayes and Moulton, 2009a). In a trial

evaluating the effectiveness of supervised tooth brushing at daycare, it would be logistically difficult and resource intensive to administer both intervention to children on an individual basis within each centre.

2. Reduce bias: In certain situations, the internal validity of a trial is optimized when treatments are randomly allocated to clusters rather than to individuals. In a trial evaluating the anti-caries effectiveness of a toothpaste, all children within a family (the cluster) may be randomized to receive either toothpaste A or toothpaste B. It would be possible to randomize individually in this instance, but compliance is likely to be better when all members of the family receive the same intervention rather than interventions that differ within a family. Similarly, the potential for contamination can be reduced. In the context of a CRT, contamination is the potential biasing of results when participants in the various arms of a trial come into contact, leading to an exchange of information between participants and possible dilution (underestimation) of the effect of the active intervention if participants in the control group benefit from the active treatment (Puffer *et al.*, 2003). Interventions evaluating oral hygiene methods have been delivered in a clustered design in a similar manner to other educational interventions (Lawrence *et al.*, 2008; Blinkhorn, 2010; Harrison *et al.*, 2010).

3. Clinical trials where multiple outcome data are recorded for each participant may be considered as clustered. Such scenarios are common in dentistry where the individual patient is the cluster, contributing a number of observations, such as multiple measurements on the same patient or multiple teeth on which the treatment is applied. For example, in the assessment of the effectiveness of two different antibiotic regimens A or B in children with one or more abscessed teeth, the child (cluster) is randomized to receive either antibiotic A or B, generating multiple observations (abscessed teeth) for the outcome within the same patient.
4. A patient's mouth, jaw, quadrant, or repeated measurements on the same patient can be considered as a cluster because they include a collection of related observations (individual teeth or a series of measurements). Furthermore, in certain occasions, especially in orthodontics, matching may be utilized as it is often feasible for different quadrants (clusters) within the same patients to receive the interventions under investigation (split mouth design). In this situation, the quadrant is the cluster, which contains several tooth units. The matching has the benefit of decreasing the required sample size compared with a CRT in which each patient receives only one intervention on all of his teeth. The smaller sample size requirement in the matched design stems from the decreased variance, which results from patients acting as their own controls (Hayes and Moulton, 2009b,c). However, matching requires availability of similar units within patients who will receive the different interventions, something that is usually difficult in some fields of dentistry (Lesaffre *et al.*, 2009).

Clustered trials require specific design, analysis, and reporting in order for the methods to be credible and the results valid (Kerry and Bland, 1998a, 1998b, 1998c; Campbell *et al.*, 2004). The value of an observation within a cluster will be more similar to the value of an observation in the same cluster than a different cluster. It is this absence of independence among observations in the same cluster or the within-cluster correlation, which creates methodological challenges. As observations within clusters are more similar, each observation within a cluster provides less information compared with an observation in a non-clustered trial. Therefore, CRTs require larger sample sizes compared with individual randomized trials, and statistical analyses that account for the fact that observations within clusters are more similar (Campbell and Grimshaw, 1998; Kerry and Bland, 1998c; Hayes and Bennett, 1999; Donner and Klar, 2000; Hayes and Moulton, 2009d). Applying methods for individually randomized trials to CRTs is inappropriate and can lead to incorrect inferences (Campbell and Grimshaw, 1998; Murray *et al.*, 2004; Varnell *et al.*, 2004; Hayes and Moulton, 2009e).

The purpose of this educational article is to explain the areas that need special attention when designing clustered trials in orthodontics and to illustrate the problems that may arise if clustered observations are erroneously treated and analysed as independent. This topic will be approached with the use of two relevant examples.

Example #1: Orthodontic bracket failure after bonding with conventional phosphoric acid etching versus self-etching primer in adolescent patients over the course of orthodontic therapy adapted from Pandis *et al.*, 2006

In this clinical study, the authors compared the risk of failure of brackets bonded with a self-etching adhesive and conventional phosphoric acid in patients followed for 12 months of active treatment. This study included multiple observations per patient, and in this design, there are clearly clustering effects, which were not accounted for during the analysis. We will use this study as a starting point and utilize simulated data in order to show the effect on the results of treating the collected data as independent (incorrect approach) or as clustered data (the correct approach).

Sample size calculations

Sample size calculations for individual randomized trials are appropriately adjusted to account for clustering (Kerry and Bland, 1998c; Hayes and Bennett, 1999; Eldridge *et al.*, 2006). The degree of variability between clusters can be given by one of two measures, the coefficient of variation (k) or the intracluster correlation coefficient (expressed as ICC or ρ or p). In the context of CRTs, both the coefficient of variation and the ICC are ways of measuring the same thing, i.e. the degree of clustering in the data.

- Coefficient of variation (k): This is the ratio of the data's standard deviation to the mean (or the proportion or the rate) of the cluster level outcomes. As the value of the standard deviation can be greater than the mean, values of k can exceed 1.
- ICC: This is a measure of the within-cluster correlation and is defined as the proportion of the total variation (between and within clusters), which can be attributed to the variation between clusters. The value of the ICC can range from -1 to $+1$, though values in the 0 to $+1$ range are usually observed. An ICC of 0 indicates that there is no within-cluster similarity in measurements, i.e. that the observations within a cluster can be considered to be independent; an ICC of 1 indicates perfect correlation. Even small ICC values are important. In CRTs, an increase in the sample size is required in order to compensate for loss of information (loss of precision/power) due to the correlated nature of the data. The required

sample size must be increased by a factor related to the degree of correlation or similarity of the outcome within clusters (Kerry and Bland, 1998c; Killip *et al.*, 2004; Eldridge *et al.*, 2006).

The ICC can be used in the calculation of the design effect to determine the effect of clustering on the sample size of a trial. This is the factor by which the sample size of a trial without clustering effects must be increased in order to give equivalent power for a cluster randomized design.

$$D = 1 + (m - 1)\rho.$$

Formula 1. Design effect (Hayes and Bennett, 1999) as a function of cluster size and ICC, where m = number of units per cluster and ρ = ICC. For example, number of m teeth contributed by each patient (cluster)

The larger the ICC (with m being the same), the larger the design effect and the required sample size for the CRT compared with the individually randomized trial with similar power (Figure 1).

In the example trial evaluating the proportion of failures for the two etching methods, if we erroneously ignore clustering of teeth within a patient and consider bond failures within patients as independent events, 800 teeth (40 patients) would have almost 80% power to detect a 5% difference in the proportion of failures (5% failure acid etching, 10% failures self-etching). If we take into account the clustering of on average 20 teeth within each patient and randomize patients to either acid etching or self-etching, assuming an ICC of 0.01, the design effect is

$$D = 1 + (m - 1)\rho = 1 + (20 - 1) * .01 = 1.19.$$

Increasing the original sample size by the design effect results in a sample size of 952 teeth, approximately 50 patients, assuming that all patients/clusters contribute 20 teeth. This is almost a 20% increase in the number of patients in order to maintain the same level of significance and power. The ICC value of 0.01 is merely used for illustrative purposes and to indicate that even small clustering effects are sufficient to have an important impact on the size of the trial. Figure 1 shows the dramatic effect of the ICC on the design effect and consequently the required sample size. Unfortunately, ICC is rarely reported in the literature and especially in dental literature. A recent study in orthodontics indicated that clustering effects are considered in the analysis in only a quarter of the studies, which have clustering due to their designs (Koletsis *et al.*, 2012). If clustering is not identified as a problem in orthodontics, then suboptimal reporting of ICC values is to be expected.

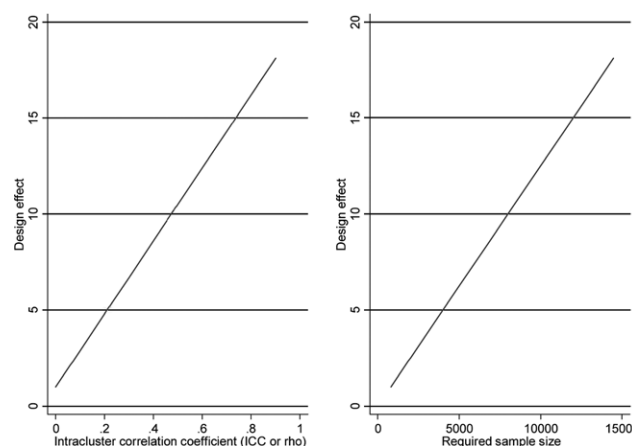


Figure 1 The effect of the intraclass correlation coefficient (ICC or rho) on the design effect and consequently on the required sample size. The sample size shows number of teeth and the assumption was that each patient contributed 20 teeth.

Randomization

The number of clusters available for randomization may be extensive or small. Where the number of clusters available for study is limited, methods such as restricted randomization, matching, and stratification may be employed in order to achieve balance in baseline characteristics between treatment arms and adjusted analyses can address unavoidable imbalances of baseline characteristics (Campbell and Grimshaw, 1998). In CRTs in orthodontics, matching can be easily and successfully achieved when all interventions (per jaw or quadrant) are applied within each patient, resulting in more precise estimates and greater study power. In our working example, cluster randomizing to acid etching or self-etching can be applied in three ways:

- Randomize patients to one etching method, conventional etching or self-etching. Each patient (cluster) contributes outcomes from 20 teeth (observations). Each patient receives one intervention.
- Randomize conventional etching to one jaw and self-etching to the other jaw. Each jaw (cluster) contributes outcomes for 10 teeth (observations). Each patient receives both interventions with the benefit of decreased variability achieved through matching. Etching method is randomly not systematically allocated to the maxilla or mandible.
- Randomize each one of the four quadrants (clusters), each contributing five teeth (observations). The advantages of matching, as stated above, are evident because both interventions are delivered for each patient. Allocation of etching method is randomly allocated to quadrants.

In our example, patients are randomized to receive one etching method either conventional or self-etching.

Analysis

CRTs can be analysed at the cluster level, taking the cluster as the unit of analysis, or at the unit level accounting for clustering, where analysis is carried out on the observations within a cluster. Analysis at a cluster level is based on the calculation of a summary value per cluster followed by simple statistical tests to compare the effect estimate between treatment arms (Hayes and Moulton, 2009a). For variable cluster size, a weighted average for proportions or odds, when binary outcomes are utilized, and weighted *t*-test are available. As the analysis is based on cluster summaries, there is some loss of information; analysis at the unit (tooth) level where the clustered nature of the data is taken into account is generally preferred, providing the number of clusters is reasonable (more than 15–20 per arm).

Analyses using the unit (in our example the unit is the tooth) are most commonly undertaken using regression models that adjust for clustering such as robust standard errors, generalized estimating equations (GEE), and random effects (Mollison, 2000). In these approaches, analysis is carried out at the unit level (tooth level) taking into account to a lesser or greater degree the clustering present in the data.

Analysis of the etching method data (simulated data) was undertaken using a variety of different methods, to illustrate the similarities and differences in results, and how inferences made from the analyses can differ. To recap, 50 patients (clusters) were randomly allocated to receive bonding with either conventional etching or self-etching on 20 units-teeth. The outcome of interest was the proportion of failures with each of the etching methods, and it was measured at tooth level. A null hypothesis of no difference in the proportions or odds of bracket failure between etching methods was specified.

The number of bond failures by etching method at the tooth level is presented in Table 1; however, this does not indicate the number of failures per cluster. The simulated data set included failures that were highly clustered in 21 out of 50 patients (Figure 2). All failures were concentrated on 11 patients for conventional etching and 10 for self-etching, with the other 29 patients having no failures at all.

The results of the various analytical approaches are given in Table 2. The result of treating the teeth as independent observations with a sample size of 500 teeth per adhesive group generates small standard errors and consequently small *P*-values and statistically significant results [odds ratio (OR) = 0.61, Chi-squared = 5.68, *df* = 1, *P* = 0.02; unconditional logistic regression OR = 0.61, 95% confidence interval (CI): 0.40, 0.92, *P* = 0.02]. Failing to take into account the clustering inherent in the data, treating the teeth as independent observations and carrying out an analysis treating bond failures within patients as independent, has resulted in a standard error smaller than it should be, leading to an increased value of the test statistic and CIs that

Table 1 Example #1: Frequency of bond failure per etching method at tooth level.

Adhesive			
Outcome	A	B	Total
No failure	436	459	895
Failure	64	41	105
	500	500	1000

are too narrow, with a real possibility of incorrect inferences being made. The fact that failures are concentrated to only a portion of patients is masked and disregarded because all observations (one per tooth) are treated as independent.

Analyses taking into account the clustering of teeth within a patient give different results. As the contribution of each individual observation has decreased relative to independent observations, so has the effective sample size. The standard errors are now greater and consequently a smaller test statistic and larger *P*-values. This phenomenon explains the fact that when the correlated data were analysed as if it were uncorrelated (no clustering present in the data), the results were significant, and when correctly treated as correlated (accounting for teeth belonging to the same patient) are no longer statistically significant (Hayes and Moulton, 2009d). Cluster level analysis gives non-significant results (OR = 0.58, 95% CI: 0.15, 2.14, *P* = 0.18). Logistic regression with robust standard errors ignores clustering when generating the effect estimates but not when estimating standard errors, and so produces the same effect estimate as unconditional logistic regression, but a larger standard error and statistically non-significant results (OR = 0.61, 95%

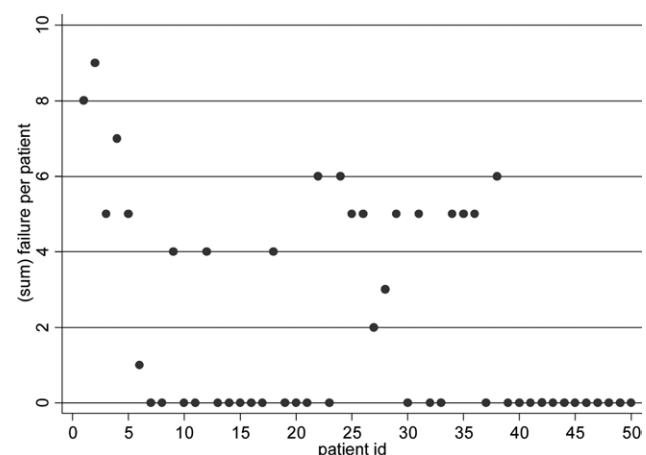


Figure 2 Scatterplot showing clustering of failures within a few patients. The x-axis shows the patient id and the y-axis indicates the number of bond failures per patient.

Table 2 Example #1: Analyses for bond failure by etching method.

	Effect estimate (odds ratio)	95% Confidence interval	Standard error	ICC	P-value	
Not accounting for correlated data						
Chi-squared	0.61	n/a	n/a	n/a	0.02	Significant
Logistic regression	0.61	0.40, 0.92	0.13	n/a	0.02	Significant
Accounting for correlated data						
Cluster level	0.58	0.15, 2.14*	0.66*	n/a	0.18**	Non-significant
Logistic regression robust SEs	0.61	0.27, 1.36	0.25	n/a	0.23	Non-significant
GEE logistic regression	0.61	0.27, 1.36	0.25	n/a	0.23	Non-significant
Random effects logistic regression	0.48	0.11, 2.11	0.36	0.6	0.33	Non-significant

GEE, generalized estimating equation; ICC, intraclass correlation coefficient; SE, standard error.

*Assuming ICC = 0.6 calculated from random effects logistic regression and using formula described by Donner and Klar (Donner and Klar, 2000).

**From two-sided *t*-test.

CI: 0.27, 1.36, $P = 0.23$). Unlike the previous approaches, GEE logistic regression, with exchangeable correlation matrix, with robust standard errors takes account of clustering in generating both the estimates and the standard errors (OR = 0.61, 95% CI: 0.27, 1.36, $P = 0.23$). In this particular example, the results using logistic regression with robust standard errors gives the same results with GEE logistic regression; however, this is not always the case and it also depends on the chosen correlation matrix. The random effects logistic regression model incorporates the variation between clusters in the likelihood. Clustering is accounted for in the effect estimates and standard errors (OR = 0.48, 95% CI: 0.11, 2.11, $P = 0.33$). From all the analyses accounting for clustering, it is inferred that there is insufficient evidence to reject the null hypothesis of no difference in proportion of failures according to etching method.

The above problem, where different analyses produce different results in terms of significance, is exacerbated when interpretation is based solely on *P*-values, and conclusions are reduced to significant or non-significant disregarding the clinical relevance of the treatment effects (Savitz, 1993; Chia, 1997; Rothman *et al.*, 2008). A study published by Petracci *et al.* on assessing bond failures using survival methods for clustered observations is a good example of how this type of data can be analysed. (Petracci *et al.*, 2009).

Example 2. Is there evidence of an effect of extraction of primary maxillary canines versus non-extraction on permanent maxillary canine impaction resolution?

The null hypothesis is that there is no difference on the resolution of canine impaction between extraction and non-extraction of primary canines. The number of primary impacted canines in each treatment group along with the success/failure for permanent canine resolution is shown in Table 3. There were 23 successful impaction resolutions (77%) in the extraction group compared with 15 successful

impaction resolutions (50%) in the non-extraction group. In the patients with two impacted canines, clustering effects are expected. This is a simulated data set consisting of 46 patients with 60 impacted canines and with a higher prevalence of successful outcomes among the patients with two impacted canines. Figure 3 shows the number of impacted canines per patient.

The results of the various analytical approaches are given in Table 4. Results of after using chi-square testing and unconditional logistic regression assuming independent observations give statistically significant results. As it was pointed out in this simulated data set, several patients have two impacted canines, and that most patients with two impacted canines have a successful or unsuccessful outcome for both. In other words, success or failure is clustered on patients with two canines. If we analyse the data using methods that account for clustering, our inferences are likely to be different compared with analyses that do not account for clustering. Failing to account for the clustering in the data leads us to the conclusion that there is evidence of beneficial effect of the intervention (OR = 3.29, chi-squared = 4.52, *df* = 1, $P = 0.03$; unconditional logistic regression OR = 3.29, 95% CI: 1.08, 9.95, $P = 0.03$); analyzing the data appropriately leads us to the opposite conclusion (cluster level analysis RR = 1.50, 95% CI: 0.99, 2.6, $P = 0.12$, robust standard errors logistic regression: OR = 3.29, 95% CI: 0.99, 10.83, $P = 0.05$, GEE logistic

Table 3 Example 2: Frequency of successful/unsuccessful resolution of canine impaction among treatment and control groups.

Intervention			
Outcome	Control	Extraction	Total
No resolution of impaction	15	7	22
Resolution of impaction	15	23	38
	30	30	60

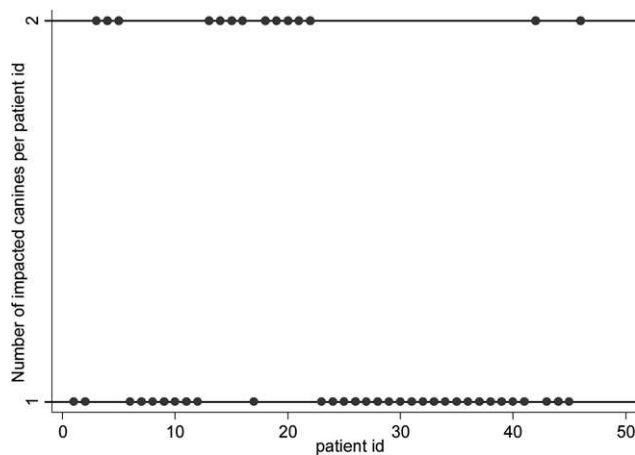


Figure 3 Scatterplot showing the patients with one or two impacted canines. Clustering effects in the presence of two impacted canines in a patient.

regression OR = 2.93, 95% CI: 0.93, 9.24, $P = 0.05$). On the basis of the analyses accounting for clustering, there is insufficient evidence to reject the null hypothesis of no difference in the resolution of canine impaction between the extraction and non-extraction groups. In reality, a 27% difference between treatment groups is large and clinically important, and it is likely that a larger sample size might have given statistically significant also for the analyses that account for clustering effects. However, the point for this example is to show how application of incorrect methods when analyzing clustered data may lead to incorrect inferences and associations that are not genuine.

The examples above used binary outcome; however, appropriate statistical methods that account for clustering are available for continuous outcomes and for rates (Hayes and Moulton, 2009b). It is worth noting that a sensitivity analysis study comparing analytical methods in CRTs that consider the clustering effect has found lower sensitivity

when the assessed outcomes are continuous compared with binary (Donner, 1982).

Reporting of CRTs

Reporting of study design and results of randomized controlled trials has been explicitly described in the CONSORT (CONsolidated Standards Of Reporting Trials) guidelines (Moher et al., 2010). The CONSORT group has also published an extension to the CONSORT guidelines specifically for CRTs (Campbell et al., 2004) in which key recommendations are to report:

- rationale for choosing a cluster randomized clinical trial
- how clustering was incorporated into design and sample size calculations
- method of randomizing the assignment such as blocking, stratification, and matching
- flow of clusters from randomization to analysis
- how the effects of clustering were incorporated into the analysis
- reporting of the ICC and the coefficient of variation

Summary

- The usual assumptions of independence of the observations are often violated in orthodontic and dental research, due to the use of clustered data such as multiple observations within patients.
- In the presence of correlated (non-independent) data, appropriate methods for randomization, sample size calculation, statistical analysis, and reporting should be followed.
- CRTs require larger sample sizes than individually randomized trials.
- Not accounting for clustering and analyzing clustered data using methods for individually randomized trials may result in incorrect inferences.
- Transparent reporting of CRTs is required.

Table 4 Example 2: Analyses of canine impaction by treatment.

	Effect estimate (odds ratio)	95% Confidence interval	Standard error	ICC	P-value	
Not accounting for correlated data						
Chi-square	3.3	n/a	n/a	n/a	0.03	Significant
Logistic regression	3.3	1.1, 9.9	1.86	n/a	0.03	Significant
Accounting for correlated data						
Cluster level analysis	1.5*	0.9, 2.6	n/a	n/a	0.12	Non-significant
Logistic regression robust SEs	3.3	0.99, 10.8	2.00	n/a	0.05	Non-significant
GEE logistic regression	2.9	0.93, 9.24	1.72	n/a	0.07	Non-significant
Random effects logistic regression**	n/a			n/a		

GEE, generalized estimating equation; ICC, intraclass correlation coefficient; SE, standard error.

*Risk ratio.

**Random effects logistic regression is not applicable for this particular example as the particular data set does not allow correct model fitting.

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